# Leveraging national forest inventory data to estimate forest carbon density status and trends for small areas

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### National Forest Inventory Data

- National Forest Inventory (NFI) programs provide critical information on forest health, sustainable management, and ecosystem change.
- Users require higher spatial and temporal resolution forest status and change parameter estimates.
- Design-based estimates are limited to large spatial and temporal scales.



### Small Area Estimation

- Small area estimation (SAE) methods have gained attention for estimating forest parameters in data-sparse settings.
  - employ statistical models to relate forest response variables to auxiliary data.
  - improve accuracy and precision over design-based approaches.



### Fay-Herriot Model

- The Fay-Herriot (FH) model is widely used in SAE applications for NFI data.
  - ▶ fit to small area direct estimates.
  - ▶ does not require exact plot locations.
- Direct estimates are often missing when sample sizes are too small or measurements are homogeneous.



# Bayesian Spatio-Temporal SAE Model

We propose a Bayesian spatio-temporal SAE model of live forest carbon density (LFCD) that

- directly uses NFI plot-level measurements,
- incorporates auxiliary covariates,
- accommodates spatially and temporally varying regression coefficients,
- appropriately quantifies uncertainty.

### Data

- We have 593,368 United States Forest Service Forest Inventory and Analysis (FIA) plot measurements collected across 3,108 counties in the CONUS from 2008 to 2021.
  - ► Exact plot locations are unknown, but plot measurements may be assigned to counties.
- We leverage remotely sensed percent tree canopy cover (TCC) as a covariate.
  - ▶ Averaged among all pixels within given county and year.

### Notation

Let

- $j = 1, \ldots, J$  index counties,
- $t = 1, \ldots, T$  index discrete years,
- $i = 1, \ldots, n_{j,t}$  index FIA plots measured in county j in year t,
  - Note, we may have  $n_{j,t} = 0$  for some j and t.
- $y_{i,j,t}$  be the LFCD (Mg/ha) at FIA plot *i* in county *j* in year *t*,
- $\mu_{j,t}$  be the latent (unobserved) mean LFCD for county j in year t,
- $\mathbf{x}_{j,t} = (1, x_{1,j,t}, \dots x_{P,j,t})^{\mathrm{T}}$  be the length P + 1 vector of covarities associated with county j in year t,
- $\tilde{\mathbf{x}}_{j,t} = (\tilde{x}_{1,j,t}, \dots \tilde{x}_{Q,j,t})^{\mathrm{T}}$  be the length  $Q \subseteq P$  vector of covariates with space-varying impact on  $\mu_{j,t}$ .

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### Model

For county j in year t, the proposed model is then

$$y_{i,j,t} = \underbrace{\mathbf{x}_{j,t}^{\mathrm{T}} \boldsymbol{\beta}_t + \tilde{\mathbf{x}}_{j,t}^{\mathrm{T}} \boldsymbol{\eta}_j + u_{j,t}}_{\mu_{j,t}} + \varepsilon_{i,j,t}, \quad i = 1, \dots, n_{j,t}, \quad (1)$$

where

- $\varepsilon_{i,j,t} \stackrel{\text{ind}}{\sim} N(0,\sigma_t^2),$
- $\beta_t$  is a length P + 1 vector of temporally-varying regression coefficients,
- $\eta_j$  is a length Q vector of space-varying regression coefficients,
- $u_{j,t}$  is a dynamically evolving spatio-temporal intercept term.

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### Temporally-varying Regression Coefficients

 $\boldsymbol{\beta}_t$  is modeled dynamically as

$$\boldsymbol{\beta}_t = \boldsymbol{\beta}_{t-1} + \boldsymbol{\xi}_t, \text{with} \tag{2}$$

$$\boldsymbol{\xi}_{t} \stackrel{\text{ind}}{\sim} MVN\left(\boldsymbol{0}, \boldsymbol{\Sigma}_{\boldsymbol{\xi}}\right), \quad t = 1, \dots, T,$$
(3)

which allows the effect of covariates in  $\mathbf{x}_{j,t}$  to have time-varying impact on the response according to the covariance structure in  $\Sigma_{\xi}$ .

### Space-varying Regression Coefficients

Writing  $\boldsymbol{\eta}_j = (\eta_{1,j}, \dots, \eta_{Q,j})^{\mathrm{T}}$  and collecting  $\boldsymbol{\eta}_q^* = (\eta_{q,1}, \dots, \eta_{q,J})^{\mathrm{T}}$ , we model  $\boldsymbol{\eta}_q^*$  as a conditional autoregressive (CAR) random effect,

$$\boldsymbol{\eta}_{q}^{*} \sim MVN\left(\mathbf{0}, \tau_{\eta,q}^{2}\mathbf{Q}(\rho_{\eta,q})\right), \quad q = 1, \dots, Q,$$
(4)

where

- $\tau_{\eta,q}^2$  is a scalar variance parameter,
- $\rho_{\eta,q}$  is a spatial correlation parameter  $(0 < \rho_{\eta,q} < 1)$ ,
- $\mathbf{Q}(\rho_{\eta,q})$  is a  $J \times J$  correlation matrix reflecting the county neighborhood structure. (See Banerjee et al. (2004) for details).

### Dynamic Spatio-temporal Intercept

 $u_{j,t}$  is modeled as a dynamically evolving CAR spatial random effect,

$$u_{j,t} = u_{j,t-1} + \omega_{j,t},\tag{5}$$

where  $u_{j,0} \equiv 0$  for all j.

Then, collecting all  $\omega_{j,t}$  for time t as  $\boldsymbol{\omega}_t = (\omega_{1,t}, \dots, \omega_{J,t})^{\mathrm{T}}$ , we specify a CAR spatial structure for  $\boldsymbol{\omega}_t$  as

$$\boldsymbol{\omega}_t \sim MVN\left(\mathbf{0}, \tau^2_{\boldsymbol{\omega},t}\mathbf{Q}(\rho_{\boldsymbol{\omega}})\right).$$
 (6)

### Direct Estimates

• Traditionally, NFI-derived quantities of interest have been estimated using design-based direct estimates.

Specifically, the direct estimate mean for  $\mu_{j,t}$  is calculated as

$$\hat{\mu}_{j,t} = \frac{1}{n_{j,t}} \sum_{i=1}^{n_j t} y_{i,j,t}, \qquad (7)$$

with associated estimate variance

$$\hat{\sigma}_{j,t}^2 = \frac{1}{n_{j,t}(n_{j,t}-1)} \sum_{i=1}^{n_j t} (y_{i,j,t} - \hat{\mu}_{j,t})^2.$$
(8)



Figure 1: Posterior mean values of live forest carbon density  $(\mu_{j,t})$ .

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TCC - Estimator - Direct - Full Model

Figure 2: Posterior mean and 95% credible intervals of LFCD  $(\mu_{j,t})$  for Tuolumne County, California, compared to direct estimate means  $(\hat{\mu}_{j,t})$  and 95% confidence intervals over time. Top row displays sample sizes  $(n_{j,t})$ .



Figure 3: Significant live forest carbon density trends (Mg/ha/year).



Estimator — Direct — Full Model

Figure 4: Average measures of bias, root mean square error (RMSE), coverage percentage, and coverage interval widths for the model and direct estimator arranged according to sample size  $n_{j,t}$ .

Each point represents the mean metric value for estimating  $\mu_{j,t}$ averaged over R = 100simulated population replicates.



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